

1 **Evidence for an Integrated Bilingual Language System from Discourse Tasks in Aphasia**

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27 **Abstract**

28 **Background:**

29 The extent to which bilingual individuals represent and process their two languages within a
30 shared or partially distinct neural architecture remains a topic of ongoing debate. While both
31 parallel and divergent patterns of impairment have been reported in bilingual aphasia, such
32 findings likely reflect a spectrum of representational overlap influenced by dominance,
33 proficiency, and task demands. Critically, few studies have examined how breakdown manifests
34 across multiple levels of linguistic structure using ecologically valid, discourse-based tasks.

35 **Aims:**

36 This study investigates whether Spanish-English bilinguals with aphasia exhibit parallel or
37 dissociable patterns of impairment across their two languages, focusing on naturalistic narrative
38 production and fine-grained analysis of error types and code-switching.

39 **Methods & Procedures:**

40 Thirteen bilingual individuals with aphasia following acquired brain injury produced story
41 retellings in both languages. Speech samples were transcribed and coded for phonological,
42 morphological, syntactic, and semantic errors, and for the word type at which they occurred.
43 Code-switches were also identified and categorized. Analyses included generalized linear
44 modeling and unsupervised clustering.

45 **Outcomes & Results:**

46 While participants made more errors in their non-dominant language, the structure and
47 distribution of errors were highly similar across languages. Clustering algorithms revealed that
48 impairments were parallel across languages and did not group by language dominance. Code-
49 switching occurred more frequently from the non-dominant to the dominant language, consistent
50 with activation-based lexical selection.

51 **Conclusions:**

52 Findings support an integrated bilingual language system that spans multiple levels of linguistic
53 representation, modulated by language dominance. Naturalistic discourse tasks allow for richer
54 characterization of bilingual language breakdown and may better inform both theoretical models
55 and clinical management of bilingual aphasia.

56

57 **1 Introduction**

58 Understanding how bilingual individuals process and represent their two languages
59 remains a central question in cognitive neuroscience, especially in the context of brain injury,
60 where disruption to the language system offers a unique opportunity to infer its underlying
61 architecture. At the heart of this inquiry lies a fundamental theoretical question: do bilinguals
62 represent and process both languages within a single, unified neurocognitive architecture, or do
63 they rely on partly distinct systems? A fully integrated model assumes that once a neural
64 mechanism exists for a given linguistic operation (e.g., plural inflection), it serves that function
65 across languages, treating cross-language variation similarly to within-language variation (e.g.,
66 across registers or dialects; Blanco-Elorrieta & Caramazza, 2021). In contrast, other perspectives
67 argue that features such as formal grammatical structure (MacWhinney et al., 2005;
68 MacWhinney, 2022), age of acquisition (Singleton, 2005; Singleton & Leśniewska, 2021), and
69 usage context can give rise to at least partially distinct representations and processing routines.
70 These views are not mutually exclusive, and growing evidence suggests that the bilingual mind
71 exhibits both integration and separation depending on the linguistic domain and level of
72 representation (Abutalebi & Green, 2013; Van de Putte et al., 2017, 2018; Malik-Moraleda et al.,
73 2024).

74 The present study directly tests these competing accounts by examining how bilingual
75 language breakdown unfolds across multiple linguistic levels following brain injury.
76 Specifically, we investigate whether Spanish–English bilinguals with post-stroke aphasia show
77 parallel or dissociable patterns of impairment across their two languages, and whether code-
78 switching behavior provides converging evidence for an integrated or partially distinct language
79 system. By analyzing naturalistic, discourse-based production, we aim to capture how bilingual
80 language organization manifests in people with aphasia under real-world communicative
81 conditions rather than decontextualized laboratory tasks.

82 Within healthy populations, much prior work has focused on the lexico-semantic level
83 (Chee et al., 1999; Dehaene et al., 1997; Hasegawa et al., 2002; Yang et al., 2017; Molinaro et
84 al., 2024), with relatively limited attention to morphosyntactic and phonological representations
85 (Declerck et al., 2020; Ou et al., 2020). While conceptual representations are generally believed
86 to be shared across languages (Van de Putte et al., 2018; Chen et al., 2024), debates continue
87 over whether morphology, syntax, and phonology are supported by common or distinct neural
88 mechanisms. Findings vary from overlapping activation patterns for L1 and L2 (Tettamanti et al.,
89 2002; Musso et al., 2003; Sakai et al., 2004; Perani & Abutalebi, 2005; Yang et al., 2017) to
90 evidence of differential activity for L1 and L2 in distinct brain regions (Lehtonen et al., 2009;
91 Golestani et al., 2006; Ip et al., 2017), and even language-specific differentiation across regions
92 of interest (Gao et al., 2023; Meykadeh et al., 2021; see Cargnelutti et al., 2019 for meta-
93 analysis; Biondo et al., 2023 for review).

94 Yet most of this evidence is correlational, leaving open the question of causality: whether
95 the same neural substrates are not only co-activated across languages but also jointly necessary
96 for their function. Studying bilinguals with post-stroke aphasia provides a rare opportunity to test
97 this directly: lesions to shared neural resources should lead to parallel impairments in both
98 languages, whereas lesions affecting language-specific regions should yield selective deficits in
99 one language only. In this way, bilingual aphasia serves as a natural lesion model that can
100 arbitrate between integrated and partially distinct accounts of bilingual language representation.
101 By analyzing the nature and distribution of errors across linguistic levels, we can infer whether
102 both languages rely on common processing mechanisms or on separable neural subsystems.

103 In addition to linguistic level, word class provides a finer-grained window into the
104 organization of bilingual language representations. A large body of research has shown that
105 aphasia can differentially affect content words (nouns, verbs, adjectives) and function words
106 (articles, prepositions, pronouns, conjunctions), reflecting their distinct roles in language
107 production and neural representation. Content words carry rich semantic and conceptual
108 information, whereas function words primarily encode grammatical relations and structural
109 dependencies (Friederici, 2011; Pulvermüller, 2002). These categories are supported by partially
110 separable neural substrates: content words engaging temporo-parietal semantic regions and
111 function words relying more heavily on left inferior frontal and temporal areas involved in
112 morphosyntactic processing (Hagoort, 2005; Tyler et al., 2011).

113 Clinically, this distinction is also evident: agrammatic and non-fluent aphasias often
114 involve disproportionate difficulty with function words and inflectional morphology, while
115 fluent aphasias tend to spare grammatical morphemes but produce semantically empty content
116 (Caramazza & Zurif, 1976; Bastiaanse & Thompson, 2012). Because bilingual aphasia might
117 manifest either as parallel or dissociable breakdown across these lexical categories, analyzing
118 errors by word class offers a more precise diagnostic lens on the structure of the bilingual
119 language system. If both languages share a unified representational and control architecture, we
120 would expect similar function–content asymmetries in both languages; if they are partly distinct,
121 the relative vulnerability of these categories could diverge. Incorporating word class thus allows
122 us to examine not only whether the two languages break down in parallel, but how deeply that
123 parallelism extends into the grammatical versus lexical strata of language.

124 Foundational neuropsychological work on bilingual aphasia has documented both parallel
125 and differential patterns of post-stroke impairment, and these contrasting outcomes have been
126 central to the debate over whether bilingual language is supported by shared or partially distinct
127 systems. Classic reports by Paradis (2001) and by Fabbro (2001) described cases in which one
128 language was disproportionately or selectively impaired, in some instances with marked
129 dissociations at specific linguistic levels (e.g., lexical access impaired in one language but
130 preserved in the other, or grammatical morphology disrupted in only one language), and in some
131 cases with asymmetric recovery trajectories across languages over time. Such cases have been
132 interpreted as evidence that, at least under some lesion and experience profiles, the two
133 languages can rely on partly independent neural resources and/or processing routines. In contrast,
134 group-level work has mainly found broadly parallel impairment across languages in bilingual
135 persons with aphasia, such that both languages show qualitatively similar error types and
136 breakdown patterns, even when they differ in overall severity (e.g., more errors in the less
137 proficient or less frequently used language; Goral et al., 2006; Green et al., 2010; Kohnert, 2004;
138 Kiran & Iakupova, 2011). On this view, apparent between-language differences largely reflect
139 graded accessibility (frequency, dominance, pre-morbid use) rather than distinct representational
140 architectures. Together, these literatures suggest that bilingual aphasia outcomes span a
141 continuum from tightly coupled to partially dissociable, and that this continuum is influenced by
142 factors such as age of acquisition, dominance, lesion site, and post-stroke language use (Kiran et
143 al., 2013).

144 Despite the richness of this work, two key gaps remain and directly motivate the present
145 study. First, most prior studies have focused on a narrow set of tasks (e.g., picture naming,
146 translation, word-level repetition) and therefore speak primarily to lexico-semantic retrieval.
147 Much less is known about whether bilinguals show parallel versus dissociable impairment

148 profiles across multiple levels of linguistic representation (i.e., phonology, morphology, syntax,
149 and semantics) when producing connected speech.

150 Second, relatively little work has examined code-switching in spontaneous discourse as a
151 theoretically meaningful diagnostic behavior. Code-switching is often described clinically as a
152 symptom of impaired control, yet it may instead reflect adaptive recruitment of whichever lexical
153 or structural resource is most available at a given moment. Critically, this behavior provides a
154 dynamic window into the functional organization of the bilingual system and therefore bears
155 directly on the question of whether both languages share or segregate neural resources. If both
156 languages draw on shared neural substrates, switches should emerge as fluid transitions within a
157 single integrated network rather than as costly shifts between distinct systems. Under this
158 activation-based view, switches are expected to occur more frequently from the non-dominant to
159 the dominant language (i.e., the language with overall greater activation; Blanco-Elorrieta &
160 Caramazza, 2021). In contrast, partially separated or inhibitory models (Green, 1998) propose
161 that producing the non-dominant language requires active suppression of the dominant one, such
162 that failures of inhibition result in “reverse dominance,” i.e., more frequent switching into the
163 non-dominant language (Gollan et al., 2014; Li & Gollan, 2018; Fadlon et al., 2019). Examining
164 the direction and frequency of code-switches in bilingual PWA thus provides a behavioral proxy
165 for how tightly coupled or segregated their two languages are at the neural level. The current
166 study addresses these gaps by analyzing naturalistic narrative production in Spanish–English
167 bilinguals with aphasia, coding error types across multiple representational levels and examining
168 the directionality and function of spontaneous code-switches.

169 Language typology may also play a role in shaping bilingual language organization.
170 Studies comparing typologically distant language pairs (e.g., English–Mandarin, Spanish–
171 Arabic) suggest that linguistic distance can influence cross-language transfer, representational
172 overlap, and recovery profiles (Ghazi-Saidi & Ansaldo, 2017). That said, if the principles of
173 language representation are indeed universal, as posited by many neurocognitive models, then
174 typological distance alone should not necessarily result in fundamentally different neural
175 architectures. At present, no clearly articulated theory specifies how or why structural differences
176 in morphology, syntax, or phonology would causally give rise to distinct neural or cognitive
177 organizations. This remains an important area for future theoretical development. The current
178 study focuses on Spanish and English, two Indo-European languages that, while different in
179 grammatical and phonological properties, share structural similarities that may facilitate
180 integration (e.g., subject-verb-object word order, alphabetic script).

181 Crucially, regardless of typological distance, different theoretical models make distinct
182 predictions about how linguistic deficits and code-switching behaviors (i.e., alternating between
183 languages within a conversation or utterance) should appear in bilingual PWA. Integrated
184 models (e.g., McClelland & Rumelhart, 1981; Rumelhart & McClelland, 1982; Blanco-Elorrieta
185 & Caramazza, 2021) predict that an impairment at a given representational level should affect
186 both languages similarly in quality, with differences arising only in degree (for example, greater
187 impairment in the non-dominant language due to less frequent use; Nadeau, 2019, or lower
188 proficiency Peñalosa et al., 2020a). Under this view, parallel error profiles would be expected
189 across languages, even if severity varies (Jacquemot et al., 2007; Jeffries et al., 2006; Lallini et
190 al., 2007; Martin et al., 1996). Crucially, these predictions must be evaluated separately for each
191 linguistic level: it is entirely plausible that phonological representations might be integrated
192 while morphosyntactic representations remain distinct.

193 In contrast, partially separated models predict that impairments can show language-
194 specific dissociations—such as impaired phonology in one language but intact phonology in the
195 other, or distinct grammatical error profiles (MacWhinney, 2022; Paradis, 2004). Crucially, such
196 dissociations have been observed in single-case reports (e.g., Fabbro, 2001; Paradis, 2001). Thus,
197 our aim here is to test the extent to which integration holds across linguistic levels using
198 naturalistic production tasks.

199

200 In all, the hypotheses we test are as follows.

- 201 1. If representations are fully integrated, error patterns should be qualitatively identical
202 across languages, differing only in quantity based on dominance or proficiency.
- 203 2. If representations are partially separated, we should observe Language × Linguistic Level
204 interactions, with some levels showing parallel deficits and others showing dissociations.
- 205 3. If activation-based lexical selection governs bilingual production, code-switching should
206 predominantly occur from the non-dominant to dominant language.

207 In the current study, we aim to characterize bilingual aphasia by analyzing the typology of
208 language impairments and the direction and function of code-switches during narrative
209 production. Rather than relying on decontextualized tasks (e.g., picture naming), we use open-
210 ended storytelling to elicit discourse that spans multiple representational levels and allows us to
211 examine real-time language use across both languages. This approach provides an ecologically
212 valid and theoretically rich opportunity to trace patterns of impairment and code-switching
213 across specific linguistic domains. By doing so, we aim to highlight possible areas of integration
214 and dissociation, and consider how these patterns might inform models of bilingual language
215 organization.

216

217

218 **2 Methods**

219

220 **2.1 Participants and Experimental procedure**

221 We collected data from 13 English-Spanish bilinguals with aphasia (9 F, 4 M; age 26-68; $M = 49$,
222 $SD = 13.74$). All PWA were living in the United States at the time of the study. Their language
223 profile was assessed with the Language Use Questionnaire (LUQ; Kastenbaum et al., 2019; Marte
224 & Carpenter et al., 2022), a notably comprehensive test that measures daily language use, language
225 proficiency of family members, language of instruction at each level of education, language
226 exposure, language confidence and self-reported language ability in hearing, speaking, reading and
227 writing across their lifetime (see Fig. 1A top panel).

228 All PWAs were tested during the subacute and chronic phase of their stroke and completed
229 the Aphasia Bank protocol (MacWhinney et al., 2011), which aims to elicit PWA's speech by
230 asking them to: i) describe their stroke experience and a memorable event, ii) describe some
231 pictures (the Broken Window sequence, the Refused Umbrella, the Cat in the Tree), iii) tell the
232 Cinderella story, and iv) describe the steps required to make a Peanut Butter and Jelly sandwich.
233 All testing was administered by trained, bilingual Spanish-English speech-language pathologists
234 and research assistants. Each testing session was conducted entirely in one language (either
235 English or Spanish), with examiners using only the target language during task administration to
236 maintain consistent language context. Testing languages were counterbalanced across sessions,
237 and examiners were instructed to minimize verbal feedback during narrative production to avoid
238 influencing participants' spontaneous language production and code-switching behaviors. The

239 speech output of these tasks constitutes the corpus for our analyses. Ten patients had confirmed
240 left-hemisphere lesions, while three had unspecified lesion lateralization in their medical records.
241 Etiologies included ischemic stroke (n = 5), hemorrhagic stroke (n= 1), unspecified CVA (n = 5),
242 Moyamoya disease (n = 1), and brain tumor (n = 1). The Western Aphasia Battery (WAB; Kertesz,
243 2007) was additionally completed to determine the severity of the aphasia in both languages. All
244 PWAs received a WAB aphasia quotient score of ≥ 60 in both languages, indicating that they were
245 mild or moderate aphasia per WAB classification (see Table. 1). WAB testing was
246 counterbalanced across languages and sessions to control for order effects. All tests were
247 administered by a trained speech-language pathologist either in person or on zoom.
248

	Age	Time post-stroke (months)	WAB_AQ English	WAB_AQ Spanish	WAB_AQ English-Spanish
Mean	49.5	57.08	76.85	79.38	-2.54
Standard Deviation	13.18	62.78	12.22	10.10	16.63
Range	27.4 – 68.3	7 – 244.7	61.7 - 97.2	60.5 - 92.6	-23.7 - 29.1

249
250 **Table 1.** Demographics and clinical characteristics of the PWA (n = 13). Values include time post-stroke (months)
251 and Western Aphasia Battery-Revised Aphasia Quotient (WAB AQ) scores in both English and Spanish, with the
252 difference between languages (English minus Spanish).
253

254

255 **2.2 Language dominance calculation**

256 Critical to our hypothesis testing is to determine i) whether there is a *quantitative* difference across
257 languages that reflects higher resilience to damage of higher frequency elements (here understood
258 also as dominant language elements), and ii) whether a *qualitatively* different profile of errors
259 emerges between dominant and non-dominant language. Both of these require defining each
260 PWA's pre-morbid language dominance, which is non-trivial due to the varying nature of every
261 person's language experience.
262

263 Traditionally, AoA and/or proficiency have been used to operationalize language
264 dominance, but they are too one-dimensional to capture the multidimensional nature of language
265 experience. To address this, we adopted a comprehensive, data-driven approach by incorporating
266 all dimensions from the Language Use Questionnaire (LUQ; Kastenbaum et al., 2019; Marte &
267 Carpenter et al., 2022)—including language usage, exposure, proficiency of family members,
268 educational language history, confidence, and self-rated ability across modalities. We calculated
269 the difference between English and Spanish for each dimension, and then applied both K-means
270 clustering and Principal Component Analysis (PCA) to these difference scores to classify
271 participants' relative language dominance (Fig. 1B).

272 K-means clustering is an algorithm that takes multidimensional data (in our case, the data
273 from every language experience dimension for each subject) and groups it on K number of clusters
274 based on similarity across all dimensions. Here, we used the algorithm to divide the data into two
275 clusters: Spanish-dominant and English-dominant. The decision to apply a two-cluster (K=2) K-
276 means solution was theoretically motivated. Our core predictions center on asymmetries between
277 a speaker's dominant and non-dominant language, predictions that hold even when the dominance
278 is subtle. Unless a participant is exactly balanced, even small asymmetries in language experience
279 are theoretically meaningful and may shape processing outcomes. In fact, including participants

280 with marginal dominance differences provides a more stringent test of our hypotheses. If effects
281 emerge across this full range, they are less likely to be artifacts of selecting only clearly unbalanced
282 individuals and more likely to reflect generalizable patterns of bilingual processing.

283 K-means clustering grouped participants into two data-driven categories. The algorithm
284 classified 7 participants in the Spanish-dominant group and 6 in the English-dominant group, based
285 on their composite language profiles.

286 To preserve the continuity of dominance and to validate the clustering outcome, we also
287 applied PCA, which reduces multidimensional data into a continuous scale. PCA is a linear
288 dimensionality reduction technique, which takes high-dimensional data (i.e., the data from every
289 dimension in the language experience questionnaire) and transforms it into low-dimensional data
290 by minimizing redundant information across related measures. In our case, we want to obtain one
291 variable from the language experience questionnaire that reflects subjects' relative language
292 dominance. We first calculated the difference between English and Spanish in each of the language
293 experience dimensions (e.g., proficiency), and submitted the vector containing these difference
294 scores to a PCA analysis. Then, we took the first component, which is the component that best
295 explains the variance in the data (in our case, 88.86%), and we obtained a raw language dominance
296 measure by calculating the projection of all dimensions for each subject along this first component.
297 To determine the directionality of this measure (i.e. whether positive indicates English-dominant
298 or Spanish-dominant), we inspected the language profile of the two subjects who received the
299 maximum and minimum values. Finally, we scaled and recentered the raw measures to the range
300 of 0-100, such that the most Spanish-dominant measure is mapped to 0 and the most English-
301 dominant measure is mapped to 100. The value represents the relative dominance of English or
302 Spanish, based not just on Age of Acquisition or Proficiency, but rather based on the compound
303 of all the answers they provided in the comprehensive Language Use Questionnaire (LUQ;
304 Kastenbaum et al., 2019).

305 Importantly, the results across the categorical and continuous measures in our analysis were
306 fully consistent: participants with a score of 50 or more in PCA were clustered in the English-
307 dominant cluster by K-means, and those with a score of less than 50 clustered as Spanish-
308 dominant. This strong convergence across methods enhances confidence in our classification
309 scheme, while also preserving the ability to explore graded dominance effects. This dual approach
310 ensures both theoretical precision and empirical robustness in our treatment of language
311 dominance.

312

313 **2.3 Data transcription and coding**

314

315 Recordings of the testing sessions were transcribed in the Computerized Language
316 Analysis Software (CLAN; MacWhinney, 2000; see Fig. 1 for a detailed description of the
317 procedure). All recordings were independently transcribed and coded following the AphasiaBank
318 transcription guidelines by native speakers of each language. To assess inter-rater reliability, we
319 followed Koo and Li's (2016) guidelines for selecting and reporting intraclass correlation
320 coefficients (ICC). Specifically, we calculated a two-way random-effects, absolute-agreement,
321 single-measure ICC, which is appropriate when each subject (here, each transcript) is
322 independently rated by two coders randomly selected from a larger pool. ICC values were
323 computed for both transcription accuracy and error categorization across all linguistic levels
324 (phonological, morphological, syntactic, and lexico-semantic). According to Koo and Li (2016),
325 ICC values below .50 indicate poor reliability, between .50–.75 moderate, between .75–.90 good,

326 and above .90 excellent reliability. Prior to formal analysis, coders independently transcribed and
327 coded 20% of the dataset (randomly selected across participants and languages). ICCs were then
328 computed for (a) number of words transcribed, (b) number of coded errors per linguistic level, and
329 (c) classification consistency of error types. All reliability estimates exceeded the threshold for
330 *good* reliability (ICC > .80).

331 To the extent that AphasiaBank was created for English, the range of error codes did not
332 cover all the possible errors one could encounter in Spanish. Thus, we developed and expanded
333 this guide to incorporate all the types of mistakes PWAs could make in Spanish. We have made
334 the guide including all the possible error types and examples in English and Spanish publicly
335 available in OSF: https://osf.io/6mwyx/?view_only=ebbfad1c0e0148c0aad2c473c4ea51f9.

336

337 The transcribers coded each error for:

338

339 1) The linguistic level where they occurred (i.e., syntactic, semantic, phonological, and
340 morphological).

341 a. We coded as syntactic errors occasions when required components of a sentence
342 were missing (e.g., “the [dog] was barking”), duplicated e.g., “the dog dog was
343 barking”), or produced in a wrong order (e.g., “the was dog barking”).

344 b. We coded as lexico-semantic errors both when PWAs substituted a target word
345 for a semantically related one, and when they had word finding problems that
346 resulted in no production.

347 c. Phonological and morphological errors included occasions where PWAs added,
348 omitted, or produced an incorrect phoneme/morpheme.

349

350 2) The class of the target word (e.g. noun, verb, pronoun, preposition, adjective, adverb).

351 3) The type of error the PWAs made (i.e., addition, substitution, omission; see Fig. 1
352 bottom right).

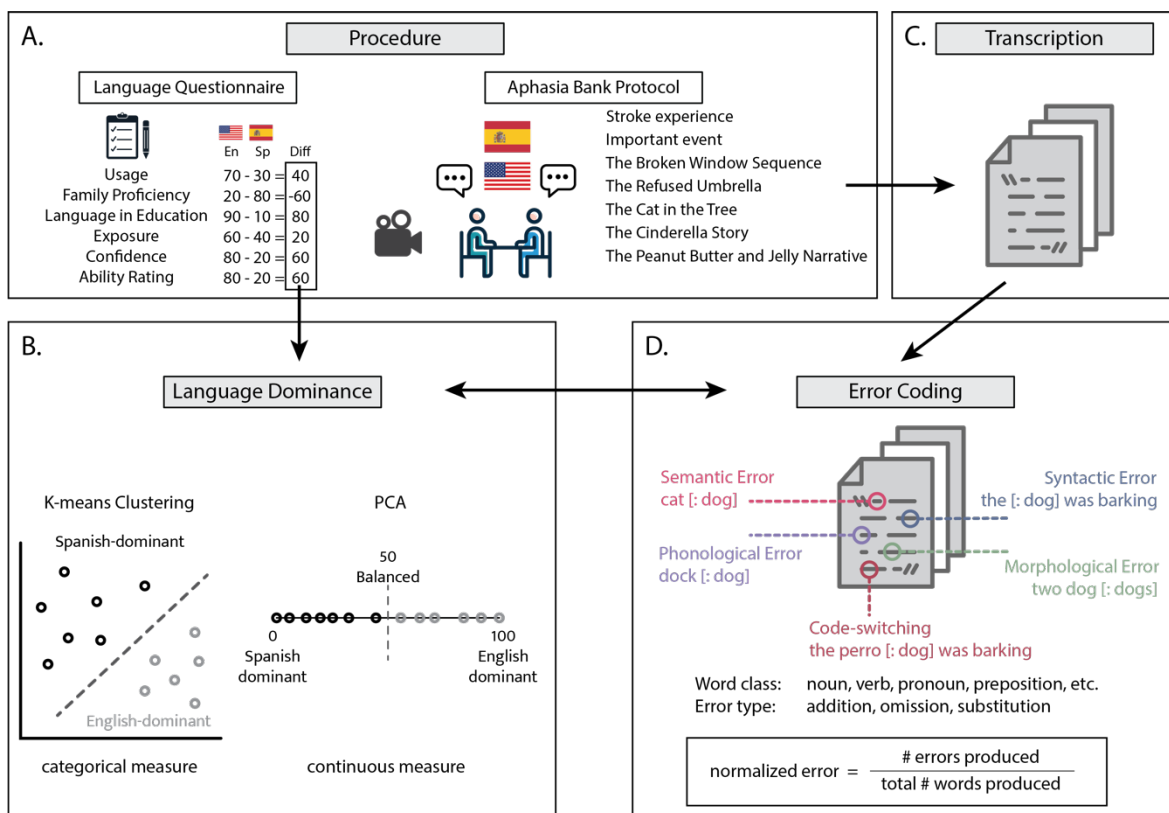


Figure 1. A. Testing and analyses procedure. During the testing session, persons with aphasia (PWA) completed the Language Use Questionnaire (LUQ; Kastenbaum et al., 2019) and the Aphasia Bank Protocol (MacWhinney et al., 2011) in English and Spanish. **B. Language dominance calculation.** The difference between languages in every measure of the language questionnaire was calculated and used to derive two measures for language dominance: a categorical measure of language dominance using K-means clustering and a continuous measure using PCA. **C. Transcription.** PWAs' speech was transcribed by two bilingual native speakers. **D. Error Coding.** PWAs' speech was coded for errors and code-switches. For each error, we coded for the level of representation (semantic, syntactic, phonological, morphological), word class (noun, verb, adjective, adverb, pronoun, determiner, preposition, conjunction), and error type (addition, omission, substitution). Normalized errors were calculated by scaling the raw frequency of errors by the total number of words produced (excluding repetition and retraction).

353
354 When a production contained multiple errors, we coded them all. For example, when a PWA
355 produced 'dafter' for the target word 'son', it was coded as a lexico-semantic error (as they tried
356 to produce the semantically related word 'daughter' for 'son') and a phonological error (for the
357 additional phoneme /f/).

358 We also identified where production switched from one language to another (i.e., there was
359 a code-switch) and coded for the class of the code-switched word (e.g., noun, verb...). When the
360 code-switch occurred for multiple-word phrases, we indicated both that the switch happened at the
361 phrase level and the class of the first word of the phrase.

362 A potential issue with using the raw number of errors/code-switches in spontaneous speech
363 production is the considerable variation in the amount of speech produced by each PWA. It is
364 crucial to address this because finding a mistake every 1000 words is significantly different from
365 finding one in every other word. Therefore, to obtain a meaningful measure of PWA's
366 impairments, it is necessary to normalize the error rates. Two common methods of normalization
367 in the literature involve dividing the raw number of errors by the total number of words or by the
368 total number of utterances. For people with aphasia, the total number of utterances is often inflated

369 by short discourse exchanges (e.g., "yeah," "mm-hmm"), so we chose to normalize by the total
370 number of words instead. Given that repetitions and reiterations common in PWA might also
371 inflate the word count, we settled for normalizing the errors using the total number of words in
372 that language excluding repetitions and retractions as the denominator. Ultimately, the three
373 measures of discourse production (total number of utterances, and total number of words with and
374 without retractions) were highly correlated ($r > 0.9$; $p < 0.001$), such that the choice of denominator
375 did not have significantly impact the normalization (Fig. 1D). We identified and excluded outliers
376 in our normalized measures using the maximum normalized residual test (Grubbs' test), which led
377 to the exclusion of one participant for the code-switching analyses.

378

379 **2.4 Statistical analyses**

380

381 To explore the impact of premorbid language dominance on post-stroke linguistic abilities,
382 we calculated the correlation between language dominance (as measured by PCA) and both (i) the
383 WAB aphasia severity score and (ii) the normalized amount of errors produced. Based on previous
384 research (Kuzmina et al., 2019; Peñaloza et al., 2020b), we expected the dominant language to
385 show milder impairment, reflected in higher WAB scores and fewer errors.

386 To examine how error rates varied as a function of linguistic and lexical factors, we fitted
387 linear mixed-effects (LME) models with normalized error rate as the dependent variable and
388 Language Dominance (dominant vs. non-dominant), Level of Representation (phonological,
389 morphological, syntactic, lexico-semantic), and Word Class (noun, verb, adjective, adverb,
390 pronoun, determiner, preposition, conjunction) as fixed effects. Subject was included as a random
391 intercept.

392

393 We followed a two-step analytic procedure:

- 394 1. Model selection: Candidate models were compared using the Bayesian Information
395 Criterion (BIC) to identify the most parsimonious model balancing explanatory power
396 and complexity (Burnham & Anderson, 2004). This procedure was used exclusively to
397 determine model structure, not to test significance.
- 398 2. Significance testing: After identifying the optimal model, we assessed the significance of
399 fixed effects and interactions within that model using Type III F-tests with Satterthwaite
400 approximation for degrees of freedom (via the *lmerTest* package).

401 The final model that minimized the BIC included the main effects of Language Dominance,
402 Level of Representation, and Word Class, as well as the Level of Representation \times Word Class
403 interaction. Adding any interaction terms involving Language Dominance increased the BIC by
404 more than 30 points, providing strong evidence that those terms did not improve model fit.

405

406 Finally, we also analyzed the effect of Language Dominance on the amount of code-
407 switching PWAs produced. To this end, we calculated the correlation between PCA language
408 dominance and the normalized amount of code-switches. We further fitted an LME model to
409 evaluate how Language Dominance and Word Class influence the normalized code-switching
410 measure (random effect: subject). Both models predict a significant effect of language dominance
411 on the amount of code-switches produced, but in opposite directions. Under an integrated language
412 system, where lexical selection occurs on the bases of highest level of activation (Blanco-Elorrieta
413 & Caramazza, 2021), code-switches should be more prevalent from the non-dominant to the
414 dominant language, because items in the dominant language have higher baseline activation and

415 are less likely to be impaired after brain damage. Contrastingly, under inhibition-based lexical
416 selection approaches, one would expect a ‘reverse dominance’ effect: more switches from the
417 dominant language to the non-dominant language (Gollan et al., 2014; Li & Gollan, 2018; Fadlon
418 et al., 2019), as a consequence of the hypothesized harder inhibition required to suppress elements
419 from the dominant as compared to the non-dominant language.

420 **3 Results**

421
422 **All effects reported below are derived from the final LME model described in Section 2.4.**
423

424 **3.1 Uneven quantitative loss across languages: dominance predicts error quantity**

425 PWAs’ post-stroke language ability was determined by their premorbid language profile as shown
426 by a significant correlation between the Language Dominance score (PCA) and i) the difference
427 in the WAB score (Table 1; $r=0.83$, $p<0.001$; Fig. 2A) and ii) the amount of errors produced in the
428 AphasiaBank tasks (Fig. 2B; $r=-0.68$, $p=0.01$). As predicted, the more dominant language was
429 more spared in our bilingual PWAs. This finding was confirmed by the LME analysis, which
430 uncovered a main effect of Language Dominance ($F=14.32$, $p<0.001$), whereby PWAs produced
431 more errors in their non-dominant language than in their dominant language across all levels of
432 representation and word classes (Figure 2C-D, Figure 3A-B).

433
434

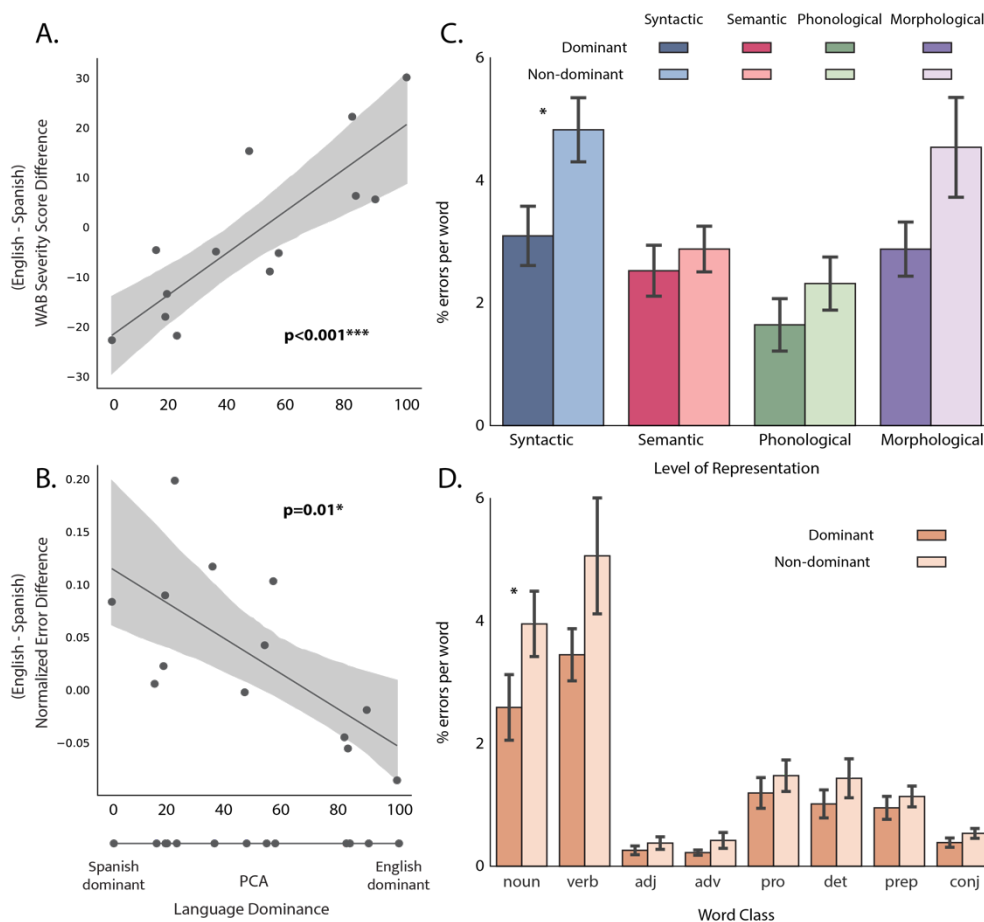


Figure 2. Effects of language dominance on post-stroke language abilities. A) shows the relationship between language dominance and difference in WAB severity score. B) shows the relationship between Language Dominance and amount of produced errors. Panels C and D show the error rates at each level of representation and word class in the dominant and non-dominant language, showing equal patterns of impairments across languages.

435

436 3.2 Parallel qualitative loss across languages: dominance does not determine error typology

437

438 To examine whether dominant and non-dominant languages show qualitatively parallel error
 439 patterns (supporting a unified language system for bilingual individuals) or whether there are
 440 selective deficits that manifest solely in one language (revealing that languages operate on
 441 independently represented systems), we focused on evaluating the interaction terms between
 442 Language Dominance, Word Class, and Level of Representation.

443 The final LME model (see Section 2.4) included the main effects of Language
 444 Dominance, Level of Representation, and Word Class, plus the Level \times Word Class interaction.
 445 Model comparison using BIC confirmed that adding either a Dominance \times Level or Dominance
 446 \times Word Class interaction worsened model fit ($\Delta\text{BIC} = +31.03$). According to conventional
 447 thresholds (Raftery, 1995), a BIC increase greater than 10 constitutes *very strong* evidence
 448 against the more complex model. In other words, the inclusion of Dominance \times Level or
 449 Dominance \times Word Class interactions did not improve the model's explanatory power
 450 sufficiently to offset their added complexity, indicating that these interactions are not supported
 451 by the data. Conceptually, this outcome is equivalent to finding that those interaction effects are
 452 not statistically significant (Burnham, & Anderson, 2004; Wagenmakers, & Farrell, 2004). In

453 short, language loss shows a parallel qualitative profile across dominant and non-dominant
454 languages (see Fig. 3A–B).

455 Within the final model, we observed a main effect of Level of Representation ($F = 11.31$,
456 $p < .001$), driven by higher error rates in morphology and syntax than in lexico-semantic or
457 phonological levels. There was also a main effect of Word Class ($F = 51.90$, $p < .001$), with most
458 errors occurring in nouns and verbs. A significant Level \times Word Class interaction ($F = 12.97$, p
459 $< .001$) reflected that different word types afford distinct opportunities for errors (e.g., semantic
460 errors are more frequent in content words, whereas function words are more constrained).

461 To further confirm that language dominance does not introduce qualitative divergence in
462 error profiles, we conducted an unsupervised K-means clustering analysis as a complementary,
463 data-driven test. We generated 26 vectors (one per participant per language), each containing 32
464 dimensions (4 representational levels \times 8 word classes), capturing the error distribution for that
465 language. We then asked whether the algorithm could recover two clusters corresponding to
466 dominance status. If dominance meaningfully shaped patients' impairment profile, the algorithm
467 should separate the data into distinct clusters for dominant versus non-dominant languages; if
468 impairment is parallel across languages, the clustering should be blind to language dominance and
469 grouping should not be based on this dimension. This logic is illustrated in Figure 3C, panels a and
470 b, which simulate the two possible outcomes: (a) distinctive clustering by language, and (b)
471 parallel distributions across languages.

472 The actual result of our clustering analysis is shown in Figure 3C, panel c. The algorithm
473 returned two clusters of similar size, but each contained a mix of dominant and non-dominant
474 language data points (Cluster A: 8 dominant / 11 non-dominant; Cluster B: 5 dominant / 2 non-
475 dominant). To quantify statistically how well the clustering aligned with the grouping expected
476 from dominance, we used the Adjusted Rand Index (ARI). The ARI measures the similarity
477 between two groupings on a scale from -1 to 1 , where 1 indicates a perfect match, 0 indicates
478 similarity no better than random chance, and negative values indicate systematic disagreement.
479 Here, we compared our grouping to the language-dominance based grouping and the observed
480 ARI was 0.02 , effectively at chance, showing that the clustering bore no relationship to what would
481 be expected if dominance shaped the error patterns.

482 To further complement the unsupervised K-means analysis, we also tested whether a
483 supervised model (i.e., one that is explicitly told which samples are dominant vs. non-dominant)
484 could learn to classify the error patterns by dominance. We used a support vector machine (SVM),
485 a standard classifier in machine learning, and evaluated it using 13-fold cross-validation so that
486 each prediction was made on data the model had not seen during training. The model's accuracy
487 was 42% , below chance level (50%) when distinguishing two groups. This result shows that even
488 when given the correct labels to learn from, the model could not find dominance-related structure
489 in the data. Together, these results support that "Language Dominance" was not a meaningful
490 dimension for grouping the data, supporting the conclusion that error distributions across
491 languages were qualitatively parallel.

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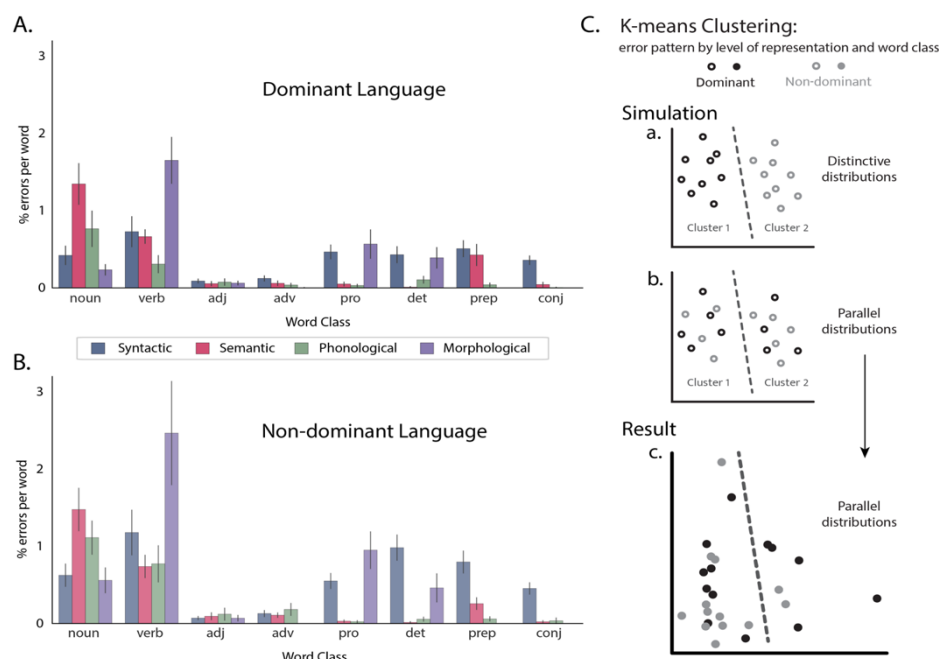


Figure 3. Effect of language dominance on error distribution. A-B: error distribution across level of representation and word class in dominant and non-dominant language. C: K-means clustering of error pattern by level of representation and word class in simulated data (a-b) and actual result (c).

493 3.3 Uneven code-switches across languages: dominance predicts language switch quantity 494 and directionality

495
496 Consistent with the prediction of an activation-based integrated model (Blanco-Elorrieta &
497 Caramazza, 2021), we observed more code-switches from the non-dominant to the dominant
498 language than vice versa (Fig. 4). We found a significant correlation between Language
499 Dominance and the difference in the normalized code-switching measure ($r=-0.64, p=0.02$) and a
500 main effect of Language Dominance in the LME model ($F=5.33, p=0.02$). There was no significant
501 main effect of Word Class ($F=1.85, p=0.08$), and we did not include interaction terms as the model
502 without them received the lowest BIC scores (-1572 vs -1549; $D = 23.75$), suggesting that
503 including interaction terms did not provide a better estimation of the data over and beyond the
504 added model complexity.

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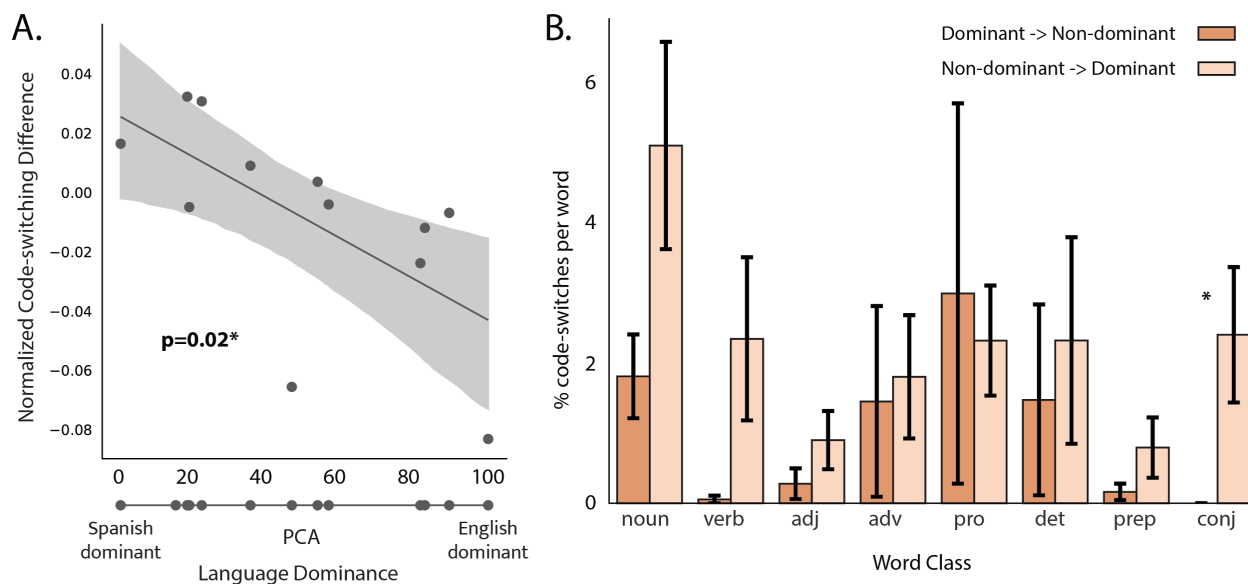


Figure 4. Effect of language dominance on the normalized code-switching measure. A: shows the correlation between PCA language dominance measure and the difference amount of code-switches. B: code-switching rate at each word class in the dominant and non-dominant language.

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4 Discussion

4.1 Implications for bilingual language organization

The present study aimed to clarify the neurocognitive architecture of bilingual language representation by analyzing error patterns and code-switching behavior in Spanish-English bilingual individuals with post-stroke aphasia (PWA). We examined how different levels of representation (phonology, morphology, syntax, and lexico-semantics) behave in bilingual language breakdown and recovery, recognizing that convergence and dissociation can co-exist within a unified yet flexible system within the context of aphasia (Abutalebi & Green, 2013; Blanco-Elorrieta & Caramazza, 2021; Van de Putte et al., 2018).

A major strength of this study lies in the use of naturalistic, open-ended speech production tasks, which allowed us to evaluate bilingual language performance across all linguistic levels and word classes in ecologically valid discourse. Unlike tightly constrained experimental paradigms, narrative production affords the opportunity to capture spontaneous linguistic behavior as it unfolds in real-time, including morphological and syntactic structures that are often excluded from item-based assessments. This approach enabled us to conduct a fine-grained analysis of error types across phonology, morphology, syntax, and lexico-semantics, levels that are rarely examined simultaneously in bilingual aphasia research.

Strikingly, across participants, we found consistent qualitative profiles of impairment in both the dominant and non-dominant language not only at every level of representation but also across different word classes. This convergence across both structural levels and lexical categories provides especially robust evidence for a shared representational system. If the two languages were supported by partially or fully distinct mechanisms, some degree of divergence would be expected in at least one linguistic domain or lexical category. Yet even when broken down into distinct word types such as content vs. function words; nouns vs. verbs the same error typologies emerged in both languages, suggesting that the underlying representations and processing principles are

537

538 largely shared across a bilingual's two languages, even after brain damage, although quantitative
539 differences in performance may emerge depending on pre-morbid proficiency and usage patterns.

540 This pattern aligns with theoretical frameworks proposing that bilinguals rely on a shared
541 representational system (e.g., Blanco-Elorrieta & Caramazza, 2021). Under this view, when
542 bilinguals experience language breakdown, the impairment reflects damage to a single integrated
543 system that supports both languages, rather than to two independent language-specific systems.
544 Consistent with this, we found that the parallelism of error types across languages persisted even
545 after controlling for language dominance and error class, and was independently confirmed
546 through supervised classification analysis and unsupervised clustering analysis. Although some
547 case reports have documented dissociations across languages (e.g., Paradis, 2001; Fabbro, 2001),
548 our findings complement rather than contradict those observations. Such dissociations may arise
549 in single cases lesion-specific factors or from differences in experiential variables such as age of
550 acquisition and proficiency asymmetries, which can differentially strengthen components of the
551 shared system. At the group level, our results indicate that shared representations constitute the
552 default organizational principle of the bilingual language system, with divergence emerging only
553 under specific individual conditions. Crucially, the presence of individual dissociations does not
554 challenge integrated models if those dissociations reflect uneven disruption or use-based
555 reinforcement within a common representational architecture, rather than separate neural
556 substrates for each language.

557 We also observed clear quantitative differences, with more errors in the non-dominant
558 language. These effects are consistent with frequency-based models of language access (Martin et
559 al., 1996; Jacquemot et al., 2007; Gollan, 2014), and suggest that availability rather than distinct
560 representational formats is the primary source of asymmetry. Future studies combining behavioral
561 data with lesion localization will be important for disambiguating these possibilities.

562 Turning to code-switching, we found that PWAs switched significantly more from the non-
563 dominant to the dominant language. This pattern is consistent with activation-based models of
564 lexical selection (Blanco-Elorrieta & Caramazza, 2021) and contrasts with inhibitory models that
565 predict more switching into the non-dominant language (Green & Abutalebi, 2013; Gollan et al.,
566 2014). Importantly, we do not claim that our data refute all inhibitory models; rather, we propose
567 that under naturalistic, self-paced conditions, lexical selection in bilingual PWAs appears to follow
568 patterns of relative availability, not inhibition.

569 Prior studies documenting reverse dominance effects often rely on forced-switch or mixed-
570 language paradigms (Li & Gollan, 2018), which engage domain-general control systems in a way
571 that narrative production does not (Fedorenko & Blank, 2020). Our results instead align with recent
572 behavioral findings in typical bilinguals and clinical populations that emphasize the
573 communicative and facilitative role of switching (Goral et al., 2019; Bihovsky et al., 2023;
574 Mooijman et al., 2024).

575 We also found a positive correlation between language proficiency and the extent of
576 switching into the dominant language, further supporting the idea that such switches reflect
577 facilitation. The study lacks a bilingual control group, which limits our ability to infer the aphasia-
578 specific impact of these effects. This limitation is unavoidable, as neurologically healthy bilinguals
579 would be unlikely to produce meaningful errors on the relatively simple set of AphasiaBank
580 narrative tasks used in this study. Despite this constraint, our results implicate activation-based
581 mechanisms as the primary force shaping bilingual lexical selection under impairment.

582
583 **4.2 Clinical implications for bilingual aphasia**

584 The present findings support a model in which bilingual individuals with aphasia share
585 underlying linguistic representations across languages, even in the presence of quantitative
586 asymmetries due to dominance. Given that neural damage typically disrupts (but does not
587 fundamentally restructure) pre-existing language systems, our results also imply that these shared
588 representations were likely present prior to stroke onset. Thus, while our data come from a
589 clinical population, they lend indirect support to models of bilingualism that posit an integrated
590 neurocognitive architecture in neurotypical individuals (e.g., Blanco-Elorrieta & Caramazza,
591 2021), that future research should test directly.

592 Clinically, this supports treatment approaches that target language-general mechanisms to
593 promote cross-language transfer (Kiran & Roberts, 2010; Scimeca et al., 2024). The observed
594 parallelism in error types across languages suggests that interventions targeting shared semantic
595 or syntactic structures in one language may facilitate improvements in the other (Marte et al.,
596 2025).

597 Code-switching behavior also has implications for diagnosis and therapy. Rather than
598 interpreting code-switches as inherently pathological, our results indicate that bilingual PWAs
599 may use switching as a strategy to maintain communicative effectiveness, consistent with the
600 notion that switching can be compensatory rather than disruptive (Blanco-Elorrieta & Pylkkänen,
601 2017; Mooijman et al., 2024). While the current study did not directly measure communicative
602 success, we interpret switching as an adaptive behavior likely grounded in activation-based
603 selection mechanisms. On this basis, clinicians may benefit from considering assessment
604 protocols that allow for responses in either language and permit switching where appropriate
605 (Goral et al., 2024). Importantly, this interpretation should not be taken as a clinical prescription.
606 Instead, we propose that clinicians consider communicative context and underlying intent when
607 evaluating switching behavior.

608 609 **4.3 Limitations and Future Directions**

610
611 While our sample size is modest, it reflects the recruitment challenges inherent to this
612 population. Excluding participants with limited discourse output introduces selection bias, as
613 these individuals likely had more severe aphasia and/or significant disparities in proficiency.
614 This exclusion may limit the generalizability of our findings to the full spectrum of bilingual
615 aphasia severity. Longitudinal studies and larger samples are essential for evaluating whether
616 recovery trajectories differ across levels of representation and between languages. Such work
617 will be critical for refining theories of bilingual language organization and informing
618 personalized, evidence-based interventions.

619 620 **5 Conclusion**

621
622 This study provides empirical support for a model of bilingual language organization in aphasia
623 that relies on shared neural and cognitive resources across languages, modulated by relative
624 language dominance. Parallel qualitative patterns of impairment across languages, together with
625 activation-based code-switching profiles, suggest that bilinguals rely on an integrated language
626 system that spans multiple levels of representation. These findings have both theoretical and
627 clinical implications, advocating for models of bilingualism that incorporate shared architecture
628 alongside usage-dependent variation, and for intervention strategies that leverage cross-language
629 commonalities to support recovery in bilingual aphasia. Because neural injury reveals rather than

630 creates underlying structures, these findings also support the inference that shared
631 representations were present in the neurotypical brain prior to injury. However, direct
632 comparisons with healthy bilinguals will be critical to fully establish which aspects of this
633 integration are universal and which may be shaped by post-stroke adaptation.

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636 **Data statement**

637 The data will be publicly available in a OSF repository.

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